



# ESTIMATION UNDER RESTRICTIONS BUILT UPON BIASED INITIAL ESTIMATORS

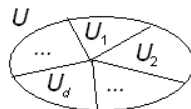
**Natalja Lepik**

Workshop of Baltic-Nordic-Ukrainian Network on Survey Statistics

24-28 August 2012



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 $d \in \mathcal{D} = \{1, 2, \dots, D\}$ .



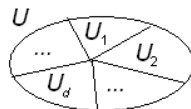
- ▶ We are interested in some domain parameters, for example domain totals,  $t^d = \sum_{i \in U_d} y_i$  with  $y_i$  being the value of study variable for object  $i$ .
- ▶ It is natural that domain totals sum up to the population total,  $\sum_{d=1}^D t^d = t = \sum_{i \in U} y_i$ .

Relationships between population parameters do not necessarily hold for the estimates in a sample!

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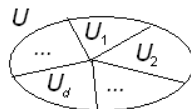
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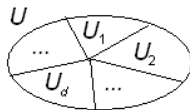
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# Cases of inconsistency



- ▶ Aggregated from registers population statistics and obtained from the sampling survey domain estimates. *Domain estimates from the survey do not sum up to the totals available from the registers.*
- ▶ The multi-survey situation: some study variables are common in two or more surveys. *Domain estimates from one survey do not sum up to the estimates of larger domains (or population totals) from another survey.*
- ▶ Domains themselves and the population total may be estimated by conceptually different estimators in the same survey. *As a result, the domain totals do not sum up to the population total, or to the relevant larger domains.*

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- ▶ It is known that any auxiliary information incorporated into estimators may increase precision of these estimators
- ▶ Known relationships between population parameters is a kind of the auxiliary information.
- ▶ If one could use this information in the estimators, one were able to make estimators more accurate and force them to satisfy desired restrictions.

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Consistency of domain estimators has been considered for some time.

This work is based on earlier results,

- ▶ the general restriction (GR) estimator (*Knottnerus 2003*),
- ▶ the GR estimator elaborated for domains (*Sõstra 2007*, *Sõstra ja Traat 2009*).

These methods produce optimal GR estimators, but they use **unbiased** or **nearly unbiased** initial estimators.

We generalize the approach by allowing **biased** initial estimators.





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# Important remarks



- ▶ All estimators and their properties are elaborated in general and can be applied for any sampling design.
- ▶ Since domain estimation is a multivariate problem, then the matrix technique is used; there are vectors of estimators under consider. The accuracy of these vectors are measured with the MSE matrices.
- ▶ The design based approach is used.
- ▶ Sample sizes in domains are assumed to be not too small, i.e. small area estimation methods are not considered here.



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Let  $\theta = (\theta_1, \dots, \theta_k)'$  be the parameter vector under study that satisfies linear restrictions:

$$\mathbf{R}\theta = 0, \quad (1)$$

where  $\mathbf{R}$  is an  $r \times k$  matrix of rank  $r$ .

- For the domain's case,

$$\mathbf{R} = (1, 1, \dots, 1, -1) : 1 \times (D + 1) \text{ and } \theta = (t_y^1, t_y^2, \dots, t_y^D, t_y)',$$

where  $D$  domain totals are  $t_y^d = \sum_{U_d} y_i$ ,  $d = 1, 2, \dots, D$ , and the population total is  $t_y = \sum_U y_i$ .



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The GR estimator for the unbiased initial estimators  $\hat{\theta} = (\hat{\theta}_1, \dots, \hat{\theta}_k)'$  (Knottnerus, 2003) that satisfies (1) has the form

$$\hat{\theta}_{GR} = (\mathbb{I} - \mathbf{K}\mathbf{R})\hat{\theta} \quad (2)$$

with

$$\text{Cov}(\hat{\theta}_{GR}) = (\mathbb{I} - \mathbf{K}\mathbf{R})\mathbf{V}, \quad (3)$$

where

$$\mathbf{K} = \mathbf{V}\mathbf{R}'(\mathbf{R}\mathbf{V}\mathbf{R}')^{-1} \quad (4)$$

and  $\mathbf{V} = \text{Cov}(\hat{\theta})$  such that  $\mathbf{R}\mathbf{V}\mathbf{R}'$  can be inverted.

The estimator  $\hat{\theta}_{GR}$  is **unbiased** and has the **minimum variance** property.



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- ▶ Knottnerus' GR estimator is **not optimal** with **biased** initial estimators.
- ▶ Three new GR estimators **handling bias**  $b = \mathbb{E}(\hat{\theta}) - \theta$  are proposed.
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## Proposition 2.2

The estimator

$$\hat{\theta}_{GR1} = (\mathbb{I} - \mathbf{K}\mathbf{R})(\hat{\theta} - \mathbf{b}), \quad (5)$$

with  $\mathbf{K} = \mathbf{V}\mathbf{R}'(\mathbf{R}\mathbf{V}\mathbf{R}')^{-1}$  is unbiased for  $\theta$ . Its variance is

$$\text{Cov}(\hat{\theta}_{GR1}) = (\mathbb{I} - \mathbf{K}\mathbf{R})\mathbf{V}, \quad (6)$$

and it is the optimal estimator among all linear estimators in  $(\hat{\theta} - \mathbf{b})$  that satisfy restriction equation (1).



## Proposition 2.3

The estimator, satisfying restrictions (1), is

$$\hat{\boldsymbol{\theta}}_{GR2} = (\mathbb{I} - \mathbf{K}^* \mathbf{R}) \hat{\boldsymbol{\theta}}, \quad (7)$$

where  $\mathbf{K}^* = \mathbf{M} \mathbf{R}' (\mathbf{R} \mathbf{M} \mathbf{R}')^{-1}$  and  $\mathbf{M} = \mathbb{M}SE(\hat{\boldsymbol{\theta}})$ . The bias of the  $\hat{\boldsymbol{\theta}}_{GR2}$  is

$$\mathbf{b}(\hat{\boldsymbol{\theta}}_{GR2}) = (\mathbb{I} - \mathbf{K}^* \mathbf{R}) \mathbf{b}, \quad (8)$$

and the mean square error matrix is

$$\mathbb{M}SE(\hat{\boldsymbol{\theta}}_{GR2}) = (\mathbb{I} - \mathbf{K}^* \mathbf{R}) \mathbf{M}. \quad (9)$$

Furthermore,  $\mathbb{M}SE(\hat{\boldsymbol{\theta}}_{GR2}) \leq \mathbf{M}$  (10)  
in the sense of Löwner ordering.



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with  $\mathbf{K}^* = \mathbf{M}\mathbf{R}'(\mathbf{R}\mathbf{M}\mathbf{R}')^{-1}$  satisfies restrictions (1) and is unbiased for  $\theta$ . It's MSE is

$$\text{MSE}(\hat{\theta}_{GR3}) = (\mathbb{I} - \mathbf{K}^* \mathbf{R})\mathbf{V}(\mathbb{I} - \mathbf{K}^* \mathbf{R})'. \quad (12)$$

Furthermore,

$$\text{MSE}(\hat{\theta}_{GR3}) \leq \mathbf{M}. \quad (13)$$



$$\hat{\boldsymbol{\theta}}_{GR1} = (\mathbb{I} - \mathbf{K}\mathbf{R})(\hat{\boldsymbol{\theta}} - \mathbf{b}),$$

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## Proposition 2.4

The mean square error matrices of the restriction estimators  $\hat{\boldsymbol{\theta}}_{GR1}$ ,  $\hat{\boldsymbol{\theta}}_{GR2}$ ,  $\hat{\boldsymbol{\theta}}_{GR3}$  and the initial estimator  $\hat{\boldsymbol{\theta}}$  can be ordered (in the sense of Löwner ordering) as following:

$$\text{MSE}(\hat{\boldsymbol{\theta}}_{GR1}) \leq \text{MSE}(\hat{\boldsymbol{\theta}}_{GR3}) \leq \text{MSE}(\hat{\boldsymbol{\theta}}_{GR2}) \leq \text{MSE}(\hat{\boldsymbol{\theta}}). \quad (14)$$



- ▶ GR estimators are general:
  - ▶ applicable for any initial estimator-vector  $\hat{\theta}$ ,
  - ▶ applicable for any restriction matrix  $\mathbf{R}$  (not only summation),
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Population:	Data of the healthcare personnel of Estonia, $N = 21764$
Study variables:	Hourly wage (continuous) Physician (binary)
Auxiliary information:	ID of healthcare institution Age Sex Education level (1..5) Domain indicator $d$ ( $d = 1, 2, 3, 4$ )



The population is divided into 4 domains by the type of the healthcare institution:

**Table:** Population and domain sizes

Domain	no. of laborers	%
1	10863	49.9
2	6742	31.0
3	3139	14.4
4	1020	4.7
Population	21764	100



- ▶ Two sample designs were applied for the population frame, SI and MN.
- ▶ For both designs 5000 independent samples were drawn with  $n = 400$ .
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# The choice of $\theta$



The restriction matrix  $\mathbf{R}$  and the restriction equation for the vector of true totals  $\theta = (\theta_1, \theta_2, \theta_3, \theta_4, \theta_P)'$ , are:

$$\mathbf{R} = (1, 1, 1, 1, -1), \mathbf{R}\theta = 0.$$

The initial estimator vector:

$$\hat{\theta} = (\hat{t}_{greg-P}^1, \hat{t}_{greg-D}^2, \hat{t}_{syn-P}^3, \hat{t}_{syn-P}^4, \hat{t}_{greg})'.$$



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# Consistency of GR, SI



Estimator	Sample	Domain $d$				Population	$\mathbf{\hat{R}\theta} =$
		$d = 1$	$d = 2$	$d = 3$	$d = 4$		
Initial	1	2450.3	1428.3	884.1	178.3	4706.2	234.9
	2	2291.8	1878.7	1003.3	191.0	5160.7	204.1
	3	2777.1	1460.7	917.8	192.5	5019.6	328.6
GR1	1	2523.2	1459.8	544.0	151.2	4678.3	0.0
	2	2382.3	1917.8	662.2	163.7	5126.0	0.0
	3	2796.2	1469.0	580.7	166.4	5012.3	0.0
GR2	1	2425.6	1417.6	707.4	165.0	4715.7	0.0
	2	2270.2	1869.4	849.8	179.5	5168.9	0.0
	3	2742.5	1445.8	670.7	173.9	5032.9	0.0
GR3	1	2463.7	1434.1	643.6	159.6	4701.0	0.0
	2	2308.4	1885.9	785.9	174.1	5154.3	0.0
	3	2780.6	1462.3	606.8	168.6	5018.2	0.0

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		$d = 1$	$d = 2$	$d = 3$	$d = 4$		
Initial	1	2450.3	1428.3	884.1	178.3	4706.2	234.9
	2	2291.8	1878.7	1003.3	191.0	5160.7	204.1
	3	2777.1	1460.7	917.8	192.5	5019.6	328.6
GR1	1	2523.2	1459.8	544.0	151.2	4678.3	0.0
	2	2382.3	1917.8	662.2	163.7	5126.0	0.0
	3	2796.2	1469.0	580.7	166.4	5012.3	0.0
GR2	1	2425.6	1417.6	707.4	165.0	4715.7	0.0
	2	2270.2	1869.4	849.8	179.5	5168.9	0.0
	3	2742.5	1445.8	670.7	173.9	5032.9	0.0
GR3	1	2463.7	1434.1	643.6	159.6	4701.0	0.0
	2	2308.4	1885.9	785.9	174.1	5154.3	0.0
	3	2780.6	1462.3	606.8	168.6	5018.2	0.0

# Consistency of GR, MN



Estimator	Sample	Domain $d$				Population	$\hat{\mathbf{R}}\hat{\theta} =$
		$d = 1$	$d = 2$	$d = 3$	$d = 4$		
Initial	1	2666.0	1676.8	952.5	178.1	4883.7	589.7
	2	2419.3	1423.5	931.1	169.9	4725.0	218.8
	3	2674.4	1241.9	908.7	164.7	4604.8	384.8
GR1	1	2533.9	1613.3	622.7	154.0	4923.8	0.0
	2	2502.3	1463.4	591.2	142.9	4699.8	0.0
	3	2661.1	1235.5	573.3	139.0	4608.8	0.0
GR2	1	2600.0	1645.1	513.6	145.0	4903.7	0.0
	2	2394.8	1411.7	768.3	157.6	4732.4	0.0
	3	2631.3	1221.1	622.4	143.1	4617.9	0.0
GR3	1	2640.5	1664.6	446.9	139.5	4891.4	0.0
	2	2435.3	1431.2	701.5	152.1	4720.1	0.0
	3	2671.8	1240.6	555.6	137.5	4605.6	0.0

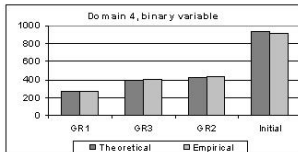
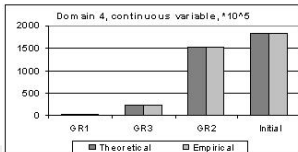
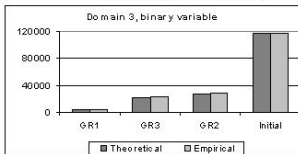
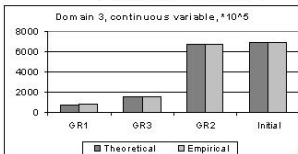
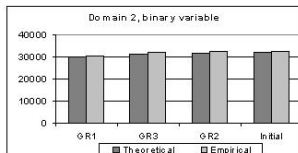
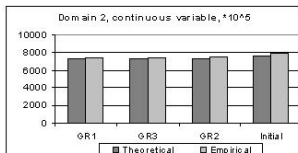
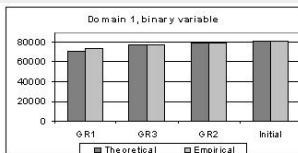
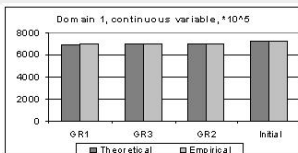
# Consistency of GR, MN



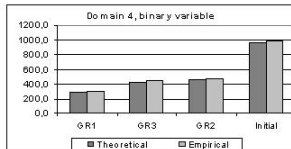
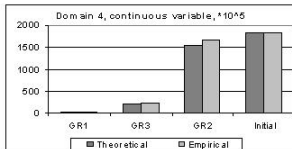
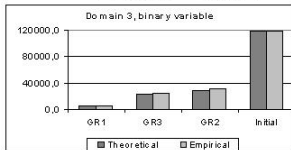
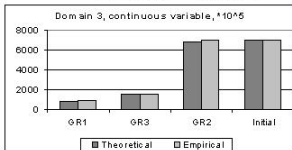
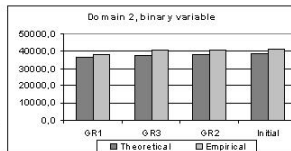
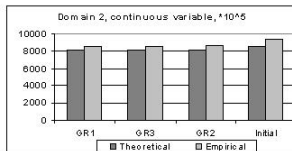
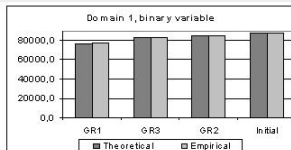
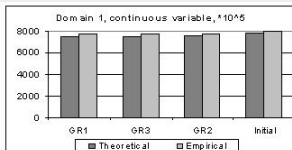
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# MSEs, SI case



# MSEs, MN case



# Final conclusion



- ▶ Three GR estimators satisfying linear restrictions, generalized to allow biased initial estimators, were proposed.
- ▶ The mean square error (MSE) and the bias expressions for the GR estimators were derived and studied.
- ▶ The ordering of the GR estimators was established, they all were more accurate than the initial estimator.



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**THANK YOU!**