# Sampling and design-based inference in finite networks 

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Finite population sampling
List-based multistage sampling:

$$
\begin{aligned}
& * \rightarrow * \sum_{\Delta}^{A} \rightarrow \Delta<! \\
& * \rightarrow *<\Delta \xrightarrow[\Delta]{\Delta} \rightarrow \Delta-\Delta \rightarrow 0
\end{aligned}
$$

NB. a special case of connections among units
"Network" $\mathcal{E}$ unconventional sampling


Key features (Zhang and Patone, 2017)

- initial sample of nodes $\mathcal{\xi}$ observation procedure by edges

- sample graph defined in terms of edges included

NB. duality of incident relationship between edge and node

Graph: $G=(\mathcal{N}, A)=$ (Nodes, edges) [digraph by default] Initial sample of nodes: $s_{1} \subset \mathcal{N} \quad\left[p\left(s_{1}\right), \pi_{i}, \pi_{i j}\right.$, etc. $]$ Observation procedure: e.g.

- induced, incident (forward, backward, reciprocal), ancestral
- snowball propagation by same procedure or adaptive

Included edges $A_{s}=A\left(s_{2}\right):$ reference set $s_{2} \subseteq \mathcal{N} \times \mathcal{N}$ e.g. induced $s_{2}=s_{1} \times s_{1}$, inc. reciprocal $s_{2}=s_{1} \times \mathcal{N} \cup \mathcal{N} \times s_{1}$

Included nodes: $\mathcal{N}_{s}=s_{1} \cup \operatorname{Inc}\left(A_{s}\right)$
Sample Graph: $G_{S}=\left(\mathcal{N}_{s}, A_{s}\right)$

Illustration: $G$ and $s_{1}=\{3,6,10\}, s_{a}=s_{1} \cup \alpha\left(s_{1}\right)$

(iii) $s_{1}=s_{a} \times s_{a}$

(iv)


## $T$-stage snowball sampling

## Initial seeds: $s_{1,0} \subset \mathcal{N}$ with successors $\alpha\left(s_{1,0}\right)$

- 1st-wave sample: $s_{1,1}=\alpha\left(s_{1,0}\right) \backslash s_{1,0}$ [seeds for 2 nd-wave]
- 2nd-wave sample: $s_{1,2}=\alpha\left(s_{1,1}\right) \backslash\left(s_{1,0} \cup s_{1,1}\right)$
- ... [ if $s_{1, t}=\emptyset$, set $\left.s_{1, t+1}=\cdots=s_{1, T}=\emptyset\right]$
- $T$-th stage sample: $s_{1, T}=\alpha\left(s_{1, T-1}\right) \backslash\left(\bigcup_{h=0}^{T-1} s_{1, h}\right)$

Sample of seeds: $s_{1}=\bigcup_{t=0}^{T-1} s_{1, t}$
I. $s_{2}=s_{1} \times \mathcal{N} \mapsto A_{s}=\bigcup_{i \in s_{1}} \bigcup_{j \in \alpha_{i}} A_{i j}$
II. $s_{2}=s_{1} \times \mathcal{N} \cup \mathcal{N} \times s_{1} \mapsto A_{s}=\bigcup_{i \in s_{1}} \bigcup_{j \in \alpha_{i}}\left(A_{i j} \cup A_{j i}\right)$

Node sample: $\mathcal{N}_{s}=s_{1} \cup \alpha\left(s_{1}\right)$

Birnbaum \& Sirken (1965): Multiplicity sampling
Example: $s_{1}$ of medical centres $(U)$, access to patients $(\Omega)$


BIG: bipartite incidence graph $G=(U, \Omega ; A)$

- bipartition $(U, \Omega)$ of $\mathcal{N}$, edges only between $U$ and $\Omega$
- e.g. $(U, \Omega)=$ (parents, children) in Lavalleè (2007)

Example: $s_{1}$ of household $(U)$, access to siblings $(\Omega)$
E.g. sampling in projection-relation graph:

- projection edges from $U$ to $P$ (persons): $\mathcal{N}=U \cup P$
- relation edges $a_{i j}=a_{j i}$ for $i, j \in P$ if $i$ and $j$ are siblings

Can use BIG with $\mathcal{N}=U \cup \Omega$ [ hypernode $k \in \Omega$ ]


Thompson (1990): Adaptive cluster sampling (ACS)


## BIG sampling

Any representation of sampling in finite graph/network

- e.g. multiplicity/indirect sampling, "network" sampling, ACS
- e.g. induced, incident, snowball sampling (Frank 1971, ..., 2011)

BIG representation $G=(U, \Omega ; A)$ for estimation

- sampling units $U$, measurement motifs $\Omega$, incidence edges $A$
- ancestral observation for design-based inference: need to know all the nodes in $U$ that could lead to the observed motifs in $\Omega_{s}$

NB. generalise the notion "multiplicity" (Birnbaum \& Sirken, 1965)

- solution: use $s_{2}^{*}=s_{1} \times s_{1}$ under $T$-stage snowball sampling
$\mathcal{C}_{q}=$ the set of all $M$ of order $q, M \subset \mathcal{N}$ and $|M|=q$
Zhang \& Patone (2017) define $q$-th order $\boldsymbol{g r a p h}$ total

$$
\theta=\sum_{M \in \mathcal{C}_{q}} y(M)
$$

Graph parameter $=$ a function of graph totals [Similarly for network totals and network parameters] Motif : a node set $M$ of specific characteristics, $M \subseteq \mathcal{N}$ NB. a motif $[M]$ may or may not have a fixed order, giving rise to graph totals with or without a given order
e.g. graph order $|\mathcal{N}|$ : 1st-order, graph size $|A|$ : 2nd-order e.g. $[M]=$ connected components, without fixed order

Example: Triads, i.e. $|M|=3$
The no. triads of size $3,2,1$, respectively, in undirected simple graph:

$$
\begin{aligned}
& \theta_{3,3}=\sum_{M \in \mathcal{C}_{3}} a_{i j} a_{j h} a_{i h} \quad[M=\{i, j, h\}] \\
& \theta_{3,2}=\sum_{M \in \mathcal{C}_{3}} a_{i j} a_{i h}\left(1-a_{j h}\right)+a_{i j} a_{j h}\left(1-a_{i h}\right)+a_{i h} a_{j h}\left(1-a_{i j}\right) \\
& \theta_{3,1}=\sum_{M \in \mathcal{C}_{3}} a_{i j}\left(1-a_{j h}\right)\left(1-a_{i h}\right)+a_{i h}\left(1-a_{i j}\right)\left(1-a_{j h}\right)+a_{j h}\left(1-a_{i j}\right)\left(1-a_{i h}\right)
\end{aligned}
$$

Relationship to the mean and variance of degrees (Frank, 1981):

$$
\begin{gathered}
\mu=\sum_{d=1}^{N} \frac{N_{d}}{N} d=\frac{2 R}{N} \quad Q=\sum_{d=1}^{N} d^{2} N_{d} \quad \sigma^{2}=\frac{Q}{N}-\mu^{2} \\
R=\frac{1}{N-2}\left(\theta_{3,1}+2 \theta_{3,2}+3 \theta_{3,3}\right) \\
Q=\frac{2}{N-1}\left(\theta_{3,1}+N \theta_{3,2}+3(N-1) \theta_{3,3}\right)
\end{gathered}
$$

BIG sampling: $\Omega=$ population set of $[M], \Omega_{s}=$ sample set of $[M]$
For convenience: enumerate the motifs as $k=1,2, \ldots$ in $\Omega$ and $\Omega_{s}$
Yhat: HT-estimator of graph total $\theta=\sum_{k \in \Omega} y_{k}$

$$
\hat{\theta}_{y}=\sum_{k \in \Omega} \delta_{k} y_{k} / \pi_{(k)}
$$

$\delta_{k}=$ inclusion indicator and $\pi_{(k)}=$ inclusion probability of motif NB. $\pi_{(k)}$ for distinction to inclusion probability $\pi_{j}$ of unit $j \in U$ NB. Under $T$-stage snowball sampling, a motif $[M]$ is observed

$$
\begin{aligned}
& \text { if } M \subseteq s_{1} \text {, where } M=\left\{i_{1}, \ldots, i_{q}\right\} \\
& \text { or if } M_{(h)} \subseteq s_{1} \text {, where } M_{(h)}=M \backslash\left\{i_{h}\right\} \text { and } 1 \leq h \leq q
\end{aligned}
$$

(Zhang and Patone, 2017)

Zhang and Patone (2017) show that

$$
\pi_{(k)}=\sum_{h=1}^{q} \operatorname{Pr}\left(M_{(h)} \subseteq s_{1}\right)-(k-1) \operatorname{Pr}\left(M \subseteq s_{1}\right)
$$

where e.g. $\operatorname{Pr}\left(M \subseteq s_{1}\right)=\pi_{\left(i_{1}\right)\left(i_{2}\right) \cdots\left(i_{q}\right)}$ is joint inclusion probability In terms of inclusion prob. in initial seed sample $s_{1,0}$, we have

$$
\pi_{\left(i_{1}\right)\left(i_{2}\right) \cdots\left(i_{q}\right)}=\sum_{L \subseteq M}(-1)^{|L|} \bar{\pi}(L)
$$

where $\bar{\pi}(L)$ is the (exclusion) probability of $L \cap s_{1}=\emptyset$ :

$$
\bar{\pi}(L)=\operatorname{Pr}\left(R_{L} \cap s_{1,0}=\emptyset\right)=\bar{\pi}_{R_{L}}=\sum_{D \subseteq R_{L}}(-1)^{|D|} \pi_{D}
$$

where $R_{L}=\bigcup_{i \in L} R_{i}$ and $R_{i}$ is the ancestors of $i$ up to the $T-1$ steps, and $\pi_{D}$ is joint inclusion probability of the nodes (in $D$ ) in $s_{1,0}$

Birnbaum and Sirken (1965): provided $\sum_{i \in U} P_{i k}=1, \forall k \in \Omega$,

$$
\theta=\sum_{k \in \Omega} y_{k}=\sum_{k \in \Omega}\left(\sum_{i \in U} P_{i k}\right) y_{k}=\sum_{i \in U}\left(\sum_{k \in \Omega} P_{i k} y_{k}\right)=\sum_{i \in U} z_{i}
$$

Zhat based on $z_{i}=\sum_{k \in \Omega} P_{i k} y_{k}$ with $P_{i k}$ 's constant of $s_{1}$ :

$$
\hat{\theta}_{z}=\sum_{i \in s_{1}} z_{i} / \pi_{i}=\sum_{i \in U} z_{i} \delta_{i} / \pi_{i}
$$

NB. Equal-share weight, given multiplicity $m_{k}=\left|A_{+k}\right|$ in BIG:

$$
P_{i k}=m_{k}^{-1} \quad \text { if }\left|A_{i k}\right|>0, \quad P_{i k}=0 \quad \text { otherwise }
$$

NB. pps-share weight: $P_{i k} \propto \pi_{i}$ if $\left|A_{i k}\right|>0, P_{i k}=0$ otherwise
NB. $\hat{\theta}_{z}$ much easier to calculate than $\hat{\theta}_{y}$ provided $m_{k}$

Example (Thompson, 1991): Two-stage ACS


|  | RRMSE (\%) |  |  |
| :---: | ---: | ---: | ---: |
|  | $\hat{s}_{1} \mid$ | $\hat{\theta}_{S C S}$ | $\hat{\theta}_{z}$ |
| 1 | $\hat{\theta}_{y}$ |  |  |
| 1 | 143.9 | 112.1 | 112.1 |
| 2 | 96.8 | 75.4 | 72.5 |
| 4 | 64.4 | 50.1 | 43.6 |
| 6 | 49.1 | 38.3 | 29.1 |
| 10 | 32.2 | 25.1 | 12.3 |

An example of graph sampling: SRS of $s_{1},\left|s_{1}\right|=3$


## An example of graph sampling: SRS of $s_{1},\left|s_{1}\right|=3$

| \# Triad types in a directed graph (Davis \& Leinhardt, 1972) |  |  |
| :---: | :---: | :---: |
|  |  |  |
| g1 | , 003 A, B, C | empty graph |
| g2 | 012 A--+B, C | graph with a single directed age |
| g3 | $102 \mathrm{~A}+-+\mathrm{B}, \mathrm{C}$ | graph with a mutual connection between two vertices |
| g4 | 021D A+--B--+C | out-star |
| g5 | 021 U A--+B+--C | in-star |
| g6 | 021C A--+B--+C | triple, directed line |
| g7 | 111D A+--+B+--C | triple |
| g8 | 111U A+--+B--+C | triple |
| g9 | 030 T A--+B+--C, A--+C | triple and transitive |
| g10 | 030C A+--B+--C, A--+C | triple |
| g11 | 201 A+--+B+--+C | triple |
| g12 | 120D A+--B--+C, A+--+C | triple and transitive |
| g13 | 120 U A--+B+--C, $A+\cdots+C$ | triple and transitive |
| g14 | 120C $A--+B-+C, A+\cdots+C$ | triple and transitive |
| g15 | 210 A--+B+--+C, $A+-+C$ | triple and transitive |
| g16 | 300 A+--+B+--+C, $A+-+C$ | triple, complete and transitive graph |



An example of graph sampling: SRS of $s_{1},\left|s_{1}\right|=3$

| $s_{2}^{*}=s_{1} \times s_{1}, s_{2}=s_{1} \times U \cup U \times s_{1}$ |  |  |  |  |
| :--- | :--- | ---: | ---: | ---: |
|  |  | RRMSE $(\%)$ |  |  |
| Parameter | $\hat{\theta}_{y}\left(s_{2}^{*}\right)$ | $\hat{\theta}_{y}\left(s_{2}\right)$ | $\hat{\theta}_{z}^{e q}\left(s_{2}\right)$ |  |
| 1st-order | Indegree | 331.261 | 26.022 |  |
| 3rd-order | Density | 0.041 | 0.003 | 0.004 |
|  | Reciprocity | 0.118 | 0.013 | 0.016 |
|  | g6 | 333.053 | 73.600 | 81.478 |
|  | g7 | 375.735 | 96.397 | 104.520 |
|  | g8 | 540.774 | 108.593 | 116.406 |
|  | g9 | 771.335 | 149.723 | 160.095 |
|  | g10 | 540.774 | 136.630 | 142.923 |
|  | g11 | 771.335 | 172.970 | 190.091 |
|  | g12 | 1095.445 | 211.943 | 230.090 |
|  | g13 | 1095.445 | 211.943 | 230.090 |
| g14 | 540.774 | 122.138 | 131.251 |  |
| g15 | 771.335 | 172.970 | 190.091 |  |
| g16 | 1095.445 | 211.943 | 230.090 |  |
| Transitivity | 0.084 | 0.028 | 0.028 |  |

Example: Sector labour flows 2015Q1-2017Q1
$|\mathcal{N}|=263$
$|A|=31120, a_{i j} \in A$ if labour flow from $i$ to $j$
Density $=0.45$, Reciprocity $=0.73$

| $s_{2}^{*}=s_{1} \times s_{1}, s_{2}=s_{1} \times U \cup U \times s_{1}$ |  |  |  |  |  |  |
| :--- | ---: | :--- | :--- | :--- | :--- | :--- |
|  | $\operatorname{RRMSE}(\%)$ |  |  |  |  |  |
| Parameter | $\left\|s_{1}\right\|=3$ |  |  | $\left\|s_{1}\right\|=6$ |  |  |
| Indegree | 75.01 | 31.76 |  | 47.84 | 22.12 |  |
| Mutual Edges | 91.20 | 37.27 | 37.42 | 57.42 | 26.01 | 26.27 |
| Density | 75.01 | 31.76 | 31.89 | 47.84 | 22.12 | 22.34 |
| Reciprocity | 62.20 | 14.00 | 14.03 | 31.35 | 8.49 | 8.57 |

## On relative efficiency

## BIG sampling with replacement (WR)

- $p_{i}=\operatorname{Pr}\left(\delta_{i}=1\right)$ for $i \in U$
- $y_{\alpha_{i}}=y_{k}$ for $k=\alpha_{i}$ and $p_{(k)}=\sum_{i \in \beta_{k}} p_{i}=p_{\beta_{k}}$
- Hansen-Hurwitz (HH) estimators

$$
\tilde{\theta}_{z}=\frac{1}{n} \sum_{i=1}^{n} \frac{z_{i}}{p_{i}} \quad \text { and } \quad \tilde{\theta}_{y}=\frac{1}{n} \sum_{i=1}^{n} \frac{y_{\alpha_{i}}}{p_{\beta_{k}}}=\frac{1}{n} \sum_{i=1}^{n} \frac{y_{k}}{p_{(k)}}
$$

Result: $V\left(\tilde{\theta}_{z}\right) \geq V\left(\tilde{\theta}_{y}\right)$, where the equality holds if $P_{i k}=p_{(k)}^{-1} p_{i}$ for $i \in \beta_{k}$ and 0 otherwise.

NB. equal-probability $s_{1} \mapsto \tilde{\theta}_{z}$ with equal-share weights

## BIG sampling without replacement (WOR)

- $\pi_{i}=\operatorname{Pr}\left(\delta_{i}=1\right)$ and $\pi_{i j}=\operatorname{Pr}\left(\delta_{i} \delta_{j}=1\right)$ for $i, j \in U$
- $\pi_{(k)}=\operatorname{Pr}\left(\delta_{k}=1\right)$ and $\pi_{(k)(l)}=\operatorname{Pr}\left(\delta_{k} \delta_{l}=1\right)$ for $k, l \in \Omega$

Result: For HT-estimators $\hat{\theta}_{y}$ and $\hat{\theta}_{z}$ with $P_{i k} \propto \pi_{i}$,

$$
\begin{aligned}
& V\left(\hat{\theta}_{z}\right)-V\left(\hat{\theta}_{y}\right)= \\
& \quad \sum_{k \neq l \in \Omega} \sum_{k} y_{k} y_{l}\left(\sum_{i \in \beta_{k}} \sum_{j \in \beta_{l}} \frac{\pi_{i j}}{\pi_{i} \pi_{j}} P_{i k} P_{j l}-\frac{\pi_{(k)(l)}}{\pi_{(k)} \pi_{(l)}}\right)
\end{aligned}
$$

NB. cluster sampling as special case $V\left(\hat{\theta}_{z}\right)=V\left(\hat{\theta}_{y}\right)$

To explore: scope of finite network sampling theory
More observation procedures, greater scope of application
Function of network totals of definite orders: yes
e.g. density, reciprocity, transitivity, etc.
e.g. "structural equivalence" ["similarity", Pearson corr.]

Parameters based on geodesic: feasible?
e.g. "closeness" centrality: inverse of mean of invserse geodesics

Measures based on fixed-point-equation: impossible?
e.g. Katz centrality: $\mathbf{x}_{N \times 1}=\alpha A \mathbf{x}+\boldsymbol{\beta}_{N \times 1}$
e.g. "regular equivalence" btw $i, j \in \mathcal{N}: \boldsymbol{\sigma}_{N \times N}=\alpha A \boldsymbol{\sigma}+\boldsymbol{I}_{N \times N}$
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